

Political Polarization in Online News Consumption

Kiran Garimella,¹ Tim Smith,² Rebecca Weiss,² Robert West³

¹MIT, ²Mozilla, ³EPFL

garimell@mit.edu, tdsmith@mozilla.com, rweiss@mozilla.com, robert.west@epfl.ch

Abstract

Political polarization appears to be on the rise, as measured by voting behavior, general affect towards opposing partisans and their parties, and contents posted and consumed online. Research over the years has focused on the role of the Web as a driver of polarization. In order to further our understanding of the factors behind online polarization, in the present work we collect and analyze Web browsing histories of tens of thousands of users alongside careful measurements of the time spent browsing various news sources. We show that online news consumption follows a polarized pattern, where users' visits to news sources aligned with their own political leaning are substantially longer than their visits to other news sources. Next, we show that such preferences hold at the individual as well as the population level, as evidenced by the emergence of clear partisan communities of news domains from aggregated browsing patterns. Finally, we tackle the important question of the role of user choices in polarization. Are users simply following the links proffered by their Web environment, or do they exacerbate partisan polarization by intentionally pursuing like-minded news sources? To answer this question, we compare browsing patterns with the underlying hyperlink structure spanned by the considered news domains, finding strong evidence of polarization in partisan browsing habits beyond that which can be explained by the hyperlink structure of the Web.

1 Introduction

Many people—according to the Pew Research Center, one in three Americans in 2018 (Geiger 2019)—prefer to receive their news online. Market data on digital publishers further indicates that search engines and social media are dominant sources of referrals to online news content, totaling over 75% of referral traffic (Parse.ly 2016). Taken together, these numbers suggest that, for many typical news consumers, the news they see is to a large part determined by the online platforms they frequent.

At the same time, recent research (Rogowski and Sutherland 2016; Webster and Abramowitz 2017; Iyengar et al. 2019) has highlighted polarization of political beliefs among the general population, with specific attention to the relationship between the agendas of political parties and the partisan

appeal of online news organizations. As online news consumption increases and evidence of polarization among the general population accumulates, understanding the mechanisms of partisan information-seeking behaviors and the properties of online news environments grows in importance.

Not only do social media constitute an important entry point for discovering news content; endorsements of news on social media have also been identified as a factor by which people trust news sources, presumably because endorsements represent a vote in favor of the source's credibility (Messing and Westwood 2014). Search engines—another important entry point for discovering news content—are driven by algorithms that exhibit a certain degree of political polarization, and additionally, the input queries of individual users themselves are also polarized in terms of their partisan alignment (Kulshrestha et al. 2017). Polarization, thus, is the product of both individual preferences and properties of the content platforms used to seek out exposure to news.

In this paper, we examine the information-seeking behaviors of individual Web users in order to contribute to understanding the relationship between mass polarization and the consumption of news content on the Web. To address this general research question, we analyzed a novel, large-scale dataset consisting of browsing logs collected via an observational field study design.

We found evidence that supports prior research (Flaxman, Goel, and Rao 2016; Bakshy, Messing, and Adamic 2015) indicating the existence of partisan preferences towards online news exposure, whether measured as the visit share in browsing histories or as the amount of time spent actively engaging with partisan news sources. Indeed, our analysis of time spent as opposed to browsing history provides more unequivocal evidence of partisan selectivity in news sources, strongly endorsing the perspective that online news consumers pursue like-minded partisan sources for political information. We find that such polarized preferences also hold when aggregating behavior across users.

We also explore the extent to which personal preferences for news sources that cater to like-minded partisan audiences dictate polarized browsing behaviors, as opposed to the homophilic network structure of news sources on the Web. Since online news sources are more likely to contain links and references to like-minded partisan content, we consid-

ered the possibility that polarized browsing patterns could mostly be attributed to the nature of content options presented to a typical online news consumer. Indeed, one of our main contributions consists of our efforts to isolate the extent to which individual partisan preferences for political information contribute to polarization patterns in news exposure, by separating the effects of user preferences from those of the biased hyperlink network of online news content. To do so, we compared polarization patterns in two sources of data: a co-browsing graph consisting of the direct observation of political information-seeking patterns of users, and a hyperlink graph obtained from a crawl of said news content during the same study period. Our analysis provides strong evidence that users specifically prefer browsing paths that lead towards polarized content, indicating that user choices play an important role in polarization.

In summary, we make the following contributions:

1. We collected browsing histories of tens of thousands of users in order to better understand polarization in online news consumption (Section 3).
2. We provide new evidence that considering the amount of time spent on news sources typically preferred by partisans reveals a different extent of polarization than observed in the prior literature, which has mostly been based on whether a user visited a site or not (Section 4.1 and 4.2).
3. We find that such polarization is prevalent at both the individual and the population level. We show that aggregated co-browsing graphs created from co-visited domains have a clear community structure consisting of left-, center-, and right-leaning domains (Section 4.3).
4. We address the question of whether polarization of news consumption can be attributed to the biased link structure of online news networks alone, or whether users' explicit content choices contribute as well. We perform a novel comparison of two distinct data sources (co-browsing choices of users vs. the link structure of the same content) and conclude that link structure alone cannot account for the high levels of selective exposure exhibited by online news consumers (Section 5).

2 Related Work

Selective exposure in online news browsing. A popular explanation given for partisan polarization in news browsing is *selective exposure* (Klapper 1960; Stroud 2010; Sears and Freedman 1967), the observation that individuals prefer to expose themselves to information that reinforces their existing attitudes and interests. As a result, many have suggested that partisan preferences predict the ideological composition of news content that people will select for consumption (Sunstein 2009; Pariser 2011).

But the impact of partisanship on selective exposure remains an open research question, as studies are mixed as to the extent to which the general population exhibits polarized browsing habits. Much of the general population appears to consume a centrist media diet, with a long tail of more extreme partisans who exclusively visit partisan-leaning web-

sites (Nelson and Webster 2017). Flaxman, Goel, and Rao (2016) examined the Web-browsing histories of 50,000 US-based users of the Bing toolbar and highlighted that a majority had the highest density of visits on a few mainstream news outlets. Guess (2018) used data from the online Web-tracking service YouGov and concluded that there was a small group of highly partisan and active users who drove a disproportionate amount of traffic to ideologically slanted websites.

Various laboratory-based studies have illustrated how individual-level selective exposure could explain large polarization effects in news exposure. Iyengar and Hahn (2009) experimentally manipulated source labels of news sites to show preferences in sources rather than content, and showed that study participants were consistent in their selection of like-minded sources for both political and non-political information. Garrett (2009) found selective preference among users from partisan websites, but also observed that a preference for like-minded content did not imply an aversion to challenging information; while people spent time on content that reinforced their beliefs, they also spent time on content with which they disagreed.

Gentzkow and Shapiro (2011) measured polarization as the degree of ideological segregation in online and offline news exposure, as well as face-to-face interactions, and determined that the Web was no more polarized than other means of consuming political information. But in a field study similar to our panel design, Peterson, Goel, and Iyengar (2018) scrutinized online browsing histories and concluded that polarized browsing habits appeared to be 2–3 times larger than previous studies had estimated. Thus, there appears to be a lack of consensus between field and laboratory studies as to whether the Web enables greater selective exposure to political information in news content.

Polarization in social media and search engines. As news consumption moves to social media (Barthel et al. 2015), researchers have started to investigate how sharing habits of news content on social media platforms are polarized along partisan lines. Numerous researchers have focused on polarization effects in individual interactions on social media platforms (Schmidt et al. 2017; Narayanan et al. 2018; Garimella et al. 2018). For example, Raghavan, Anderson, and Kleinberg (2018) examined the structure of online interactions in order to determine how different media sources are “invoked” in replies on social media posts, concluding that polarization exists within the context of the way people use social media platforms outside of news exposure.

Existing research has, however, not reached a consensus on whether the use of social media networks and search engines increases the likelihood of polarized browsing habits (Nelson and Webster 2017; Flaxman, Goel, and Rao 2016; Nikolov et al. 2019). For example, Barberá et al. (2015) showed that polarization in news consumption on social media had been overestimated, while studies using survey data (Barthel et al. 2015) showed that people self-report polarized browsing habits. Bakshy, Messing, and Adamic (2015) compared the role of individual choices and algorithmic personalization on news consumption on Facebook

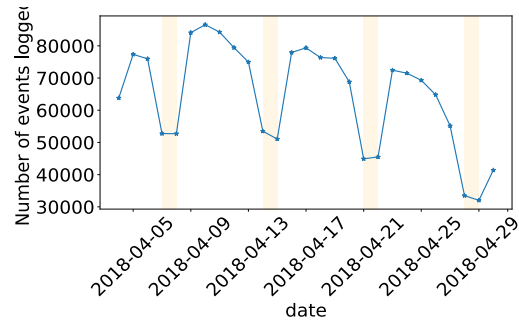
and concluded that personal preferences, rather than content recommendations on social feeds, drove content choices.

Measuring political leaning of sources and users. Social media user behavior has also been used to identify polarity (or leaning) scores of news domains and hence to quantify polarization. Bakshy, Messing, and Adamic (2015) created a list of 500 news domains and assigned a leaning score to a news domain based on the fraction of self-identified conservative users on Facebook who had shared the domain. A similar, yet much larger, list was created by Robertson et al. (2018), who used a panel of Twitter users of self-identified political leaning in order to compute leaning scores for sources. Following a similar methodology, Kulshrestha et al. (2017) used lists of Twitter users who were tagged as Democrats or Republicans and quantified the leaning of users based on whom they followed from these lists. Then, based on the following of various news organizations on Twitter, they identified the leaning of news domains. Ribeiro et al. (2018) made use of Facebook’s advertising platform data to compute the leaning of a news source by the extent to which liberals or conservatives were over- or under-represented among its audience. Finally, Lahoti, Garimella, and Gionis (2018) made use of the relationship between user preferences and their news consumption in order to jointly estimate both user leaning and news domain leaning on Twitter.

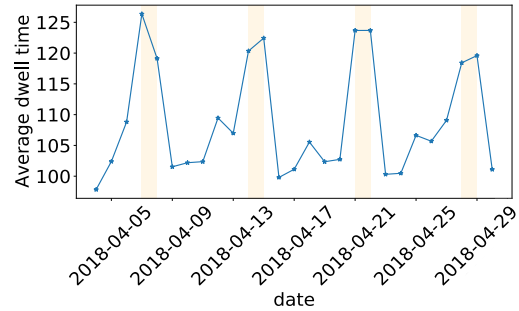
3 Data and Definitions

We randomly recruited participants from the active Firefox user population on 5 April 2018 for enrollment in an extension-based study. The sampled participants were presented with a prompt asking whether the user was interested in participating in a research study involving close scrutiny of their browsing behavior, including their browsing histories. The prompt and process of informed consent was approved by Mozilla’s legal and ethical processes. When a user consented, a browser extension was installed, which created a record logging the user’s active dwell time on content organized by URL for each day of browsing activity. Dwell time is implemented via the `nsIdleService` API of the browser, which checks an operating system-level idle detection API every five seconds. In short, this method creates a log of activity per URL that is persistent across multiple tabs and windows. This is novel compared to many other methods of measuring attention to online content, which typically measure duration as a result of analytics or tracker information visible from server logs or simply browsing history without duration measurements. If a participant opened a link but spent very little time engaging with the content of the page, browsing history-based measurements would weigh such a visit equally to a high-engagement visit.

In contrast to visit counts, the dwell time measurement consists of three distinct states that a site visit can be in: active, idle, or not in session. In our data collection, a visit to a website, either through a user opening a new window or a new tab, triggers an event for that URL to be logged. After five seconds of inactivity (measured as a lack of keyboard or mouse interaction with the browser), the site visit



(a) Number of events



(b) Average dwell time

Figure 1: (a) Daily number of events logged during the three-week study period. (b) Average dwell time per page visit (in seconds). The shaded bands indicate weekends.

is considered to have entered into an idle state. If the user resumes interaction with the content within 30 minutes, we resume our logging for the same URL. If the user moves the focus away from the URL (e.g., by opening a new window or transferring focus to a new tab) or after 30 minutes of idle time have elapsed, the session is ended.

Activity statistics of study participants. Our dataset consists of browsing logs from a panel of 24,036 unique participants, based in the US, for a period of three weeks starting on 5 April 2018. The dataset consists of 241 million distinct page visits with a median of 6,074 page visits per individual. Figure 1a shows the activity of all users. We observe that activity (number of events logged) decreases over the course of the study period, and that there is a noticeable drop during weekends. On the contrary, Figure 1b shows that the average dwell time per visit is larger on weekends than during the week, and that, even as the number of events logged decreases over the course of the study period (Figure 1a), this is not the case for the average dwell time. This suggests that, despite participant churn, the general pattern of interactions remains consistent. For robustness, we repeated all analyses on the subset of users who remained active throughout the observation period, finding no qualitative disagreement from the results obtained on the set of all users.

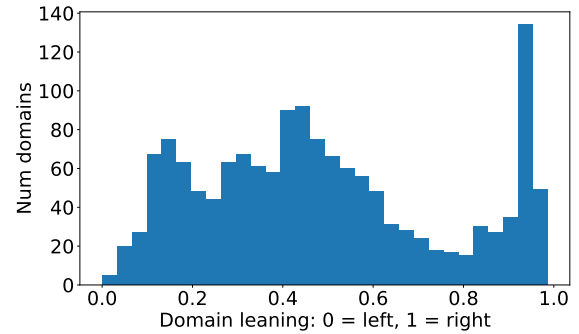
Bias and limitations of dataset. Since the data was constructed from the browsing activity of a panel of opt-in Fire-

fox users, we acknowledge that the data is likely to be biased. As we do not have demographic covariates for members of the study panel, we are not able to compare the representativity of this sample against more conventional measures of the general population, as Guess (2018) and Peterson, Goel, and Iyengar (2018) were able to do. Since it is possible that the preference to choose Firefox as a primary browser could be associated with one’s latent news preferences (in that political predispositions could predict both the likelihood to use Firefox and exposure to news content), we considered the possibility that the participants in this study might deviate from the general population in their overall news preferences. We compared the top 100 most visited news sites in our Firefox logs with the top 100 most popular (with respect to unique monthly US visitors) news sites according to Alexa.com, a prominent website ranking service. Spearman’s rank correlation between the two lists was 0.85, suggesting that our study participants were generally exposed to similar popular sources of online content as the general population.

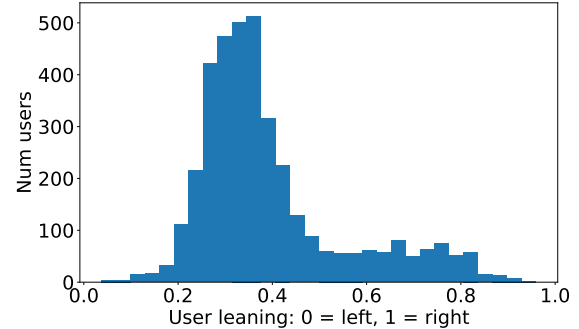
Partisan leaning computation. All our analysis is at the domain level. We obtained partisan leaning scores for all domains from Robertson et al. (2018), where “leaning” is defined as the estimate of a news outlet’s political ideological alignment with either a conservative or a liberal audience. The leaning score of a domain assumes a value between 0 and 1; it is defined as the average self-reported ideology of users who have shared pages from that domain on Twitter, where 0 corresponds to maximally liberal, and 1 to maximally conservative. The original list provided by Robertson et al. (2018) has leaning scores for over 19,000 domains, including domains such as facebook.com, google.com, and instagram.com. However, since we are specifically interested in understanding news consumption behavior, we only used a subset of news domains, obtained from the popular journalism watch-dog website mediabiasfactcheck.com. The subset contained 2,873 news domains that were still online in September 2020, and further discarding domains without a leaning score in Robertson et al.’s list resulted in a set of 1,295 news domains.

One should bear in mind that defining the leaning of a domain based on the fraction of self-identified users with a specific political leaning could be biased due to self-selection. Although we acknowledge this bias, we argue that any assignment of a leaning score to a domain (say, by a journalist or by a professional organization) will be subject to similar personal biases. The leaning scores we use have been shown to be highly correlated with the scores from Bakshy, Messing, and Adamic (2015), who used a similar methodology on Facebook, thus providing some evidence that the scores are robust. We use the list of Robertson et al. (2018) because it is larger and more up-to-date. The 1,295 studied domains include a mix of news sources (e.g., nytimes.com, washingtonpost.com), cable TV sites (e.g., msnbc.com, foxnews.com), etc. The distribution of the leaning scores of domains is shown in Figure 2a.

In this study, we only consider participants in our analysis who in total made at least 50 visits to pages from the



(a) Histogram of domain leaning scores



(b) Histogram of user leaning scores

Figure 2: Histogram of leaning scores for (a) domains and (b) users.

1,295 considered news domains. This threshold restricted the dataset to 6,575 participants. We compute the political leaning of individual participants as the weighted average leaning score of the domains they visited, where each domain was weighted by the number of times it had been visited by the respective user. The distribution of participants’ leaning is shown in Figure 2b. We can see that, on average, the studied participants are leaning liberal, as estimated through their online news browsing history. Note that, in the rest of the paper, we shall use the terms leaning/polarity/ideology, left-leaning/liberal, and right-leaning/conservative interchangeably.

Browsing graph. Important parts of our analysis are based on the bipartite (user-by-domain) *browsing graph* $G = (U, D, E)$, where $U = \{u_1, \dots, u_n\}$ are the n users, $D = \{d_1, \dots, d_m\}$ are the m news domains, and E is the set of edges $\{\{u_i, d_j\} : u_i \in U, d_j \in D\}$ between users and domains. Each edge $\{u_i, d_j\}$ is associated with a weight w_{ij} that captures the average dwell time spent by user u_i on pages from domain d_j (e.g., if a user on average spends 31 seconds on pages of nytimes.com, the corresponding weight is 31). The bipartite graph G may also be thought of as a matrix whose rows are users, whose columns are news domains, and whose entries indicate the average dwell time of a user on a domain.

Co-browsing graph. Using the bipartite browsing graph G , we performed a one-mode projection onto the news do-

	Nodes	Edges	Avg. degree	Top PageRank
Co-browsing	1,295	176,945	273	nytimes.com, washingtonpost.com, cnn.com, theguardian.com, npr.com,
Hyperlink	1,295	323,036	498	washingtonpost.com, huffingtonpost.com, nytimes.com, businessinsider.com, theatlantic.com

Table 1: Statistics of co-browsing and hyperlink graphs.

mains D by counting all paths of length 2 between two domains. We call the graph resulting from the projection the *co-browsing graph*. Edges between two domains d_1 and d_2 in the co-browsing graph are weighted in order to represent the number of users who visited both d_1 and d_2 . To count co-visits of domains, we aggregated URLs at the domain level, considered only those domains for a given user with which they had actively engaged for at least 60 seconds in total (measured via the sum of the user’s dwell times on the domain), and counted multiple visits by the same user to the same domain only once. Summary statistics of the co-browsing graph are listed in Table 1.

Hyperlink graph. In order to investigate users’ browsing patterns in the light of how online news websites link to one another, we constructed the *hyperlink graph* of the network of news domains. The Common Crawl project¹ regularly crawls the Web and releases datasets of the discovered web-pages on a regular schedule, which are generally considered the most extensive publicly available sources regarding the structure of the Web (Meusel et al. 2014). Common Crawl also periodically releases versions of the hyperlink graph of the Web aggregated at the domain level, specifying for each domain to which other domains it contains hyperlinks.² Our analysis works at the domain level and is restricted to the subgraph defined by the set of 1,295 news domains D also used in constructing the browsing and co-browsing graphs. In order to enable a meaningful comparison with the co-browsing graph, we ignore the directionality of edges and consider the hyperlink graph as an undirected graph. The hyperlink graph represents static connections between domains, and results should be interpreted in this framework. Summary statistics of the hyperlink graph are listed in Table 1. In particular, we observe that the hyperlink graph is denser than the co-browsing graph.

The browsing, co-browsing, and hyperlink graphs capture different aspects of how users browsing the Web move from domain to domain. The bipartite browsing graph reflects choices made by users; it aggregates each user’s entire browsing history. The co-browsing graph further aggregates across all users, thus giving a macro picture of the browsing

patterns of online news. In contrast, the hyperlink graph reflects the choices made by content creators to link between pages. In this sense, the hyperlink graph also represents the set of possible direct links that users *could* have taken while browsing.

4 Polarization Patterns in News Browsing

In this section, we build on top of the above-described datasets to characterize polarization in news browsing. We start by investigating dwell times spent on various domains as a way of showing the existence of polarization. Next, we analyze communities of domains in the co-browsing graph in order to show that certain groups of news domains tend to be co-visited by the same user.

4.1 Polarization in Dwell Times

First, we assess the extent of polarized browsing habits among participants by examining whether there is any clear structure in the bipartite browsing graph through the use of spectral graph co-clustering techniques (Dhillon 2001; Role, Morbieu, and Nadif 2019). Unlike graph clustering techniques, which can only cluster one dimension (users or domains) independently, co-clustering techniques find the dependencies between participants and domains. Recall that edge weights of the bipartite browsing graph capture each user’s average dwell time for each domain (the average amount of time they spent actively engaging with pages from the respective domain), such that, by co-clustering the bipartite browsing graph, we obtain clusters of similar domains visited by similar users. To avoid individual activity biases, a given user’s domain-level dwell times were z -score-standardized across all domains they had visited. Given this setup, polarized browsing habits would appear as tight clusters of domains with average domain leaning scores that significantly differ across clusters.

Running the co-clustering procedure for a number of co-clusters ranging from two to five, we found that when using three co-clusters, a liberal-leaning (left), a centrist, and a conservative-leaning (right) cluster of domains emerged. Since, moreover, using more than three co-clusters merely split one of the co-clusters into two, we decided to set the number of co-clusters to three. For each cluster of domains and for each cluster of users, we computed the distribution of their leaning scores. Figure 3a shows a box plot of the distribution of domain cluster leanings. Similar trends emerged in the leaning of user clusters (not shown).

Recall that the co-clustered browsing graph was based on dwell time information only; neither domain nor user leanings were used in the co-clustering process. In this light, Figure 3a, which shows that the cluster-specific distributions of leaning scores differ widely across domain clusters, implies that domains that are visited in similar ways by similar users tend to be similar with respect to their leaning—a clear indication of partisan polarization in browsing.

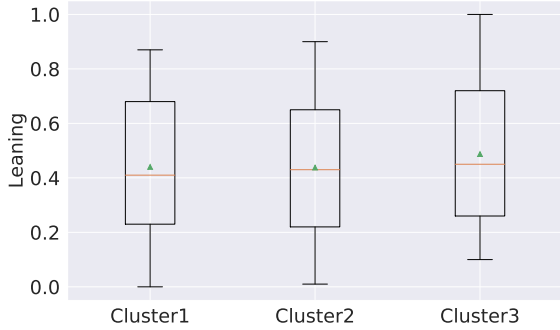
In order to analyze whether the clustering structure can be identified by website visits alone (without considering dwell times), we also performed a co-clustering on an unweighted version of the bipartite browsing graph. As observed in Figure 3b, the separation in the communities is

¹<http://commoncrawl.org/>

²e.g., <http://commoncrawl.org/2019/02/host-and-domain-level-web-graphs-nov-dec-2018-jan-2019/>



(a) Weighted browsing graph



(b) Unweighted browsing graph

Figure 3: Boxplots of leaning scores per domain cluster, obtained by co-clustering (a) the weighted bipartite (user-by-domain) browsing graph, (b) an unweighted version of the browsing graph. Green dots are means, red lines are medians, box boundaries are quartiles.

not as pronounced in the unweighted as in the weighted version, suggesting that dwell time–based measurements reveal stronger polarization patterns. Given our novel methodology to compute the dwell time of users, this is the first large-scale study to establish the existence of polarization in terms of time spent by users on partisan websites as opposed to browsing histories alone.

4.2 Time Spent “On the Other Side”

Next, we investigate whether users of different leanings engage differently with domains of different leanings. To do so, we first split the set of news domains into left-leaning (leaning below 0.4), center (leaning between 0.4 and 0.6), and right-leaning (leaning above 0.6) domains. Then, for each user, we computed their average dwell time for each domain and z -score-standardized their average dwell times across domains (such that, for a fixed user, the mean across domains is 0, with a standard deviation of 1), in order to remove effects due to the fact that some users generally spend more, and some less, time engaging with visited pages. Next, we binned users into five buckets based on their leanings and computed the mean standardized dwell time in each user bucket, separately for the three leaning-based domain groups.

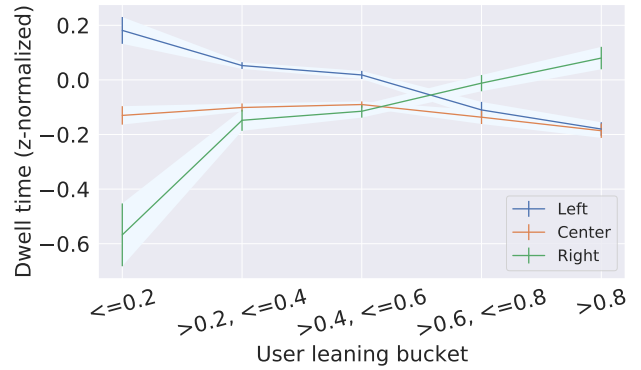


Figure 4: Average dwell time (z -score-standardized within users) spent by users of different leanings (x -axis) on domains of different leanings (three curves). Error bars show 95% confidence intervals. We see that users spend significantly more time when visiting pages on domains aligned with their own leaning.

Figure 4 shows the dwell time for left-leaning, center, and right-leaning sites for users with varying leanings. We observe that the more extreme participants on either side of the leaning spectrum dwell significantly longer when visiting like-minded online news content. For domains without a clear leaning (“center”), there is no such upward or downward trend. Although Garrett (2009) observed in a lab experiment that people spend *more*—rather than less—time on certain types of content from the other side, we emphasize that our population is several orders of magnitude larger than theirs and that our data was collected “in the wild”, outside of a lab setting.

4.3 Community Structure in Co-browsing Graph

Above, we analyzed *individual* visits to news domains by participants and found patterns of polarization in the dwell time spent upon visits of domains of various ideological alignment. Now, we consider browsing patterns *across participants* by analyzing the community structure of the co-browsing graph. The co-browsing graph represents users’ choices in what domains are co-visited, so if people only select a specific set of domains that appeal to their partisan ideology, we should observe communities of ideologically similar domains in the co-browsing graph.

To determine whether this is the case, we applied a community detection algorithm based on the Louvain method (De Meo et al. 2011) to the co-browsing graph, producing a partition of the graph nodes into communities that maximized modularity. Modularity measures the quality of a partition and is high if nodes are densely connected inside communities, and sparsely connected across communities. The method, which determines the number of communities automatically, identified three communities. Figure 5a plots the distribution of domain leanings for each of the three communities. Although domain leanings were not used in the community detection process, a clear stratification of the three communities along partisan lines emerges, with clusters 1,

Community 1	Community 2	Community 3
nytimes.com	theguardian.com	foxnews.com
cnn.com	theatlantic.com	wsj.com
nbcnews.com	thinkprogress.org	breitbart.com
latimes.com	motherjones.com	drudgereport.com
cbsnews.com	dailykos.com	dailywire.com

Table 2: Top five domains with respect to PageRank centrality in each community of the co-browsing graph.

2, and 3 containing mostly left-leaning, centrist, and right-leaning domains, respectively. Table 2 shows the five domains with the highest PageRank centrality in each of the communities.

We also used a k -core decomposition (Batagelj and Zaveršnik 2003) of the co-browsing graph to confirm this structure and found the same three strongly connected cores (communities), also separated along the ideological spectrum.

In addition to the co-browsing graph, we also ran community detection on the hyperlink graph. Here, too, three communities emerge; the results are shown in Figure 5b. Although the three communities of the hyperlink graph are qualitatively similar to those of the co-browsing graph (left, center, right), the stratification is not as clear in the hyperlink graph as in the co-browsing graph (Figure 5a): the modularity of the co-browsing graph is twice that of the hyperlink graph (0.11 vs. 0.05).

5 Selective Exposure vs. Structure of the Web

Finally, we explore whether polarization in browsing is due to selective information seeking on behalf of participants or whether it is the result of the link structure of the Web environment with which they interact.

Homophily in hyperlink and co-browsing graphs. To begin, we compute, for both the co-browsing graph and the hyperlink graph, the “neighborhood leaning” of each node (i.e., domain) d , defined as the average leaning of all nodes linked to d by an edge. (Recall that both the co-browsing graph as well as the hyperlink graph are treated as undirected graphs.) Figure 6a, computed on the co-browsing graph, plots one point per domain d , showing d ’s own leaning on the x -axis, and d ’s neighborhood leaning on the y -axis. We observe that, as we move from left to right, the neighborhood leaning increases, a clear sign of polarization in co-browsing patterns. Figure 6b, which was computed on the hyperlink graph rather than on the co-browsing graph, exposes a similar pattern, but with a significantly lower Pearson correlation coefficient (0.696 vs. 0.777, $p < 0.01$). The higher correlation coefficient for the co-browsing graph indicates that neighborhoods are more homophilic in the co-browsing graph, compared to the hyperlink graph.

As a caveat, we point out that, as seen in Table 1, the average degree of nodes in the hyperlink graph is higher than that in the co-browsing graph, which might bias the neighborhood leaning scores: a higher average degree could *a priori* lead to a more diverse set of neighbors, which could in turn lead the hyperlink graph to have less leaning homophily.

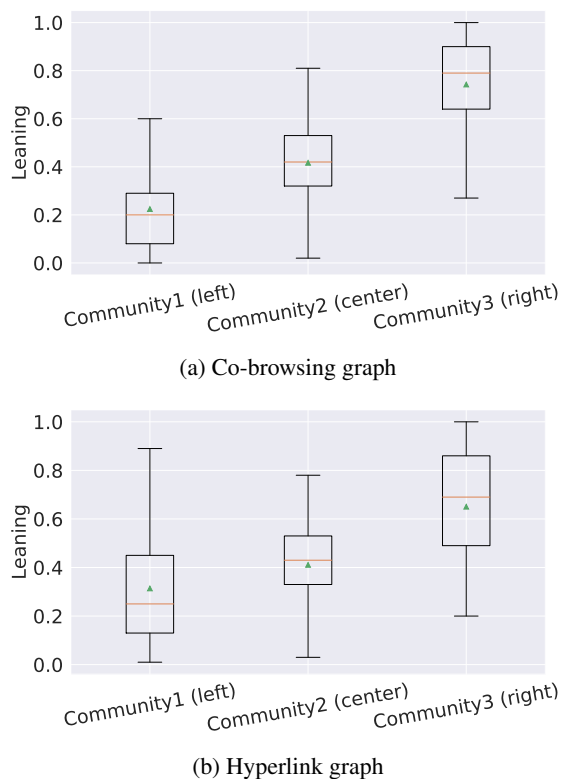
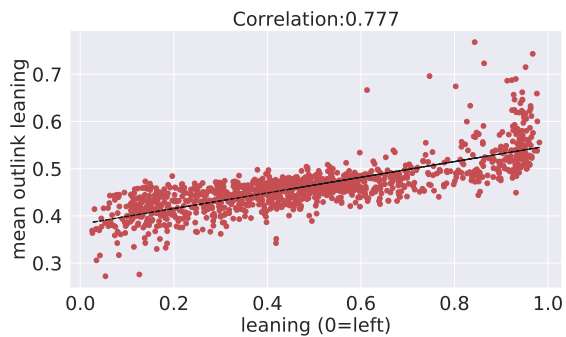


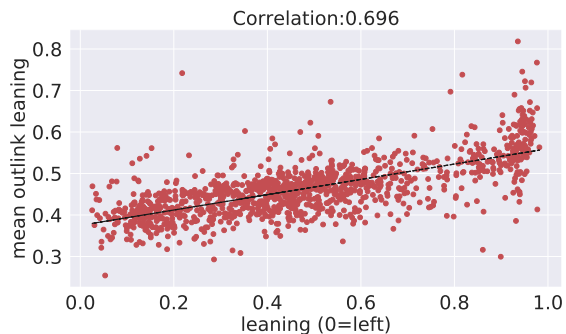
Figure 5: Boxplots of leaning scores per community of domains, obtained by running community detection on (a) the co-browsing graph, (b) the hyperlink graph. Green dots are means, red lines are medians, box boundaries are quartiles.

Presence of neighbors in hyperlink versus co-browsing graphs. Next, we compare neighborhoods of nodes in the co-browsing and hyperlink graphs. Intuitively, an edge being present in the hyperlink graph but absent from the co-browsing graph corresponds to participants deliberately avoiding the link when browsing. To investigate such situations more closely, we computed, for each node in the hyperlink graph, the absolute difference in leaning from each of its neighbors. We then ranked the neighbors by this difference and bucketed them into deciles (only considering nodes with more than 10 neighbors, which eliminated five nodes from the analysis). Next, for each decile, we computed the fraction of edges present in the co-browsing graph. Our hypothesis was that, in comparison to the hyperlink graph, edges in the co-browsing graph have more homophily, such that a larger fraction of edges would be present in the lower deciles (corresponding to neighbors in the hyperlink graph that are more similar in leaning to the focal node). Figure 7a shows that, indeed, as we move from left to right (i.e., as the leaning difference between neighbors in the hyperlink graph increases), the fraction of edges present in the co-browsing graph decreases.

The above analysis does not take into account the edge weights present in the co-browsing graph; i.e., it does not distinguish between the case where one single participant



(a) Co-browsing graph



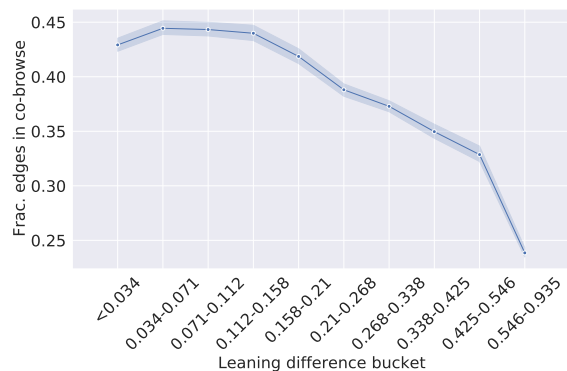
(b) Hyperlink graph

Figure 6: Leaning of each domain (x -axis) vs. average leaning of the domain’s neighbors (y -axis) in (a) the co-browsing graph and (b) the hyperlink graph.

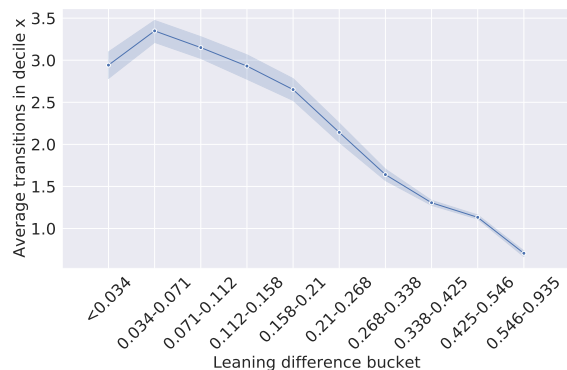
co-visited two domains from the case where a large number of participants did so. To obtain the weighted version of Figure 7a, we computed, for each decile, the average weight of the corresponding edges in the co-browsing graph, with edges absent from the co-browsing graph contributing values of zero. (Note that the previous analysis is a special case of the present analysis, with all non-zero edge weights set to the value of 1.) This way, we obtain, for each decile, the average number of users co-visiting the domains linked by the corresponding set of edges in the hyperlink graph. The results, shown in Figure 7b, are consistent with those of the unweighted analysis (Figure 7a), confirming that dissimilar (with respect to leaning) neighbors of a node are less sought out by users when browsing.

As mentioned above, a higher average degree could lead to a more diverse set of neighbors and hence, this result could be drastically different for nodes having vastly different degrees. To investigate this possibility, we also repeated the above analysis after stratifying nodes into 10 equally sized groups with respect to their degree in the hyperlink graph. Inspecting version of Figure 7 plotted separately for each of the 10 degree strata (not shown), we observed a similar pattern within each stratum, indicating the robustness of the result.

Multi-hop neighborhoods: random walks. The above analysis only considered the immediate neighborhood of a



(a) Hyperlink graph vs. unweighted co-browsing graph



(b) Hyperlink graph vs. weighted co-browsing graph

Figure 7: (a) Fraction of edges from the hyperlink graph that are also present in the co-browsing graph, when considering, for each node, only those edges that connect the node to a neighbor whose leaning difference falls into the respective decile (x -axis; deciles were computed per node over all its neighbors). (b) Analogous analysis, but with average edge weights, rather than fraction of edges, on the y -axis. The plots show that hyperlinks leading to more similar (with respect to leaning) neighbors are more likely to be chosen by users.

node when assessing whether personal preferences are more polarized than the static link structure. To understand how polarized the broader neighborhood of a node is, we conduct a random walk–based analysis, as follows. Starting from each node u , we perform a random walk and, at each step j , measure the absolute difference δ_{uj} in leaning between u and the node at step j . Then, for each step j , we compute the *neighborhood distance* δ_j at step j as the mean value of δ_{uj} over all nodes u in the graph. To ensure the robustness of the results, we repeat the random walk 100 times for each node u and use the mean δ_{uj} for each u and j . If a graph is more polarized, a node’s neighborhood will be more similar to the node itself, and the neighborhood distance δ_j will be smaller.

We ran the above procedure (with random walks of length 10) on the following three graphs: (i) the co-browsing graph without edge weights, (ii) the co-browsing graph with edge

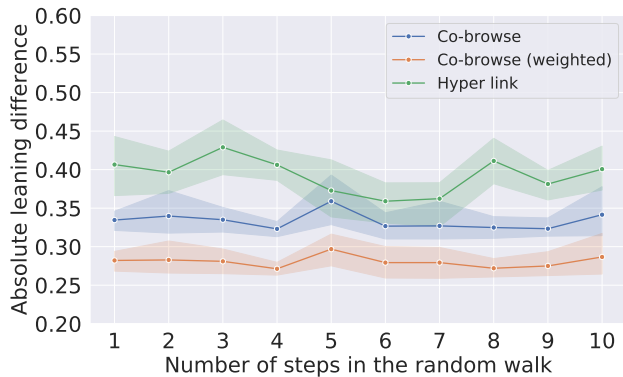


Figure 8: Leaning difference between nodes and their multi-hop neighborhoods in the weighted and unweighted co-browsing graphs and in the hyperlink graph, computed by performing random walks of length 10. Error bands show 95% confidence intervals. The weighted co-browsing graph is most homophilic, followed by the unweighted co-browsing graph and the hyperlink graph.

weights, and (iii) the hyperlink graph. Random walks in the different graphs capture different types of structure and participant choices: The weighted co-browsing graph captures the participants’ choices, biasing the random walk towards edges that were empirically taken by more participants. The unweighted co-browsing graph still takes participant choices into account to a certain extent, but makes transitions to all neighboring domains equally likely. Finally, a random walk on the hyperlink graph simulates random browsing of the considered domains.

The results are shown in Figure 8, with one curve for each of the three graphs. Each point indicates the average neighborhood distance for a given step j along the random walks. The significantly lower values of neighborhood distance for the weighted co-browsing graph (red) indicate that participants’ browsing choices are more biased towards domains with a similar leaning than what would be expected from someone randomly browsing the hyperlink graph induced by the included news domains (green). There are also significant differences in the neighborhood distance values between the *unweighted* co-browsing graph (blue) and the hyperlink graph (green), indicating inherently different (more homophilic) structure in the co-browsing graph, compared to the hyperlink graph, even after multiple hops of a random walking. Figure 8 thus strengthens the results of Figure 7 beyond the immediate neighborhood of nodes.

To summarize, the co-browsing graph appears to be significantly more polarized than the hyperlink graph, and we conclude that the polarization observed in the news consumption of the participants of our study cannot be explained by the hyperlink structure of the Web alone. Rather, participants’ explicit choices play an important role as well.

6 Conclusions

In this paper, we used browsing logs collected on the client side over the course of three weeks from a large opt-in panel of Firefox users in order to measure political polarization in

online news consumption. In various analyses, we provided evidence of pronounced polarization patterns in online news consumption.

By analyzing dwell time (the amount of time during which a user actively interacts with a webpage), we observed that users engage more deeply with news content aligned with their own leaning (estimated, by proxy, via the average leaning of the news they consume overall).

Applying a co-clustering algorithm to the bipartite browsing graph, which encodes average dwell times for all users on all visited domains, revealed three coherent, frequently co-visited groups of news domains with vastly different leaning distributions, corresponding to left-leaning, center, and right-leaning news domains, respectively—yet another indicator of polarization in news consumption.

Moving from the individual level to the population level, we contracted the bipartite user-by-domain browsing graph into a domain-by-domain co-browsing graph. Applying a community detection algorithm to the co-browsing graph gave rise to the same polarized pattern, with each cluster corresponding to one distinct political leaning.

A similarly polarized pattern—though less pronounced—was discovered by applying community detection to the hyperlink graph spanned by the news domains included in the study, rather than to the co-browsing graph. Thus, in order to investigate the possibility that polarization in news consumption might be solely due to the inherent hyperlink structure of the Web, we compared the co-browsing and hyperlink graphs in various ways, concluding that the polarized hyperlink structure alone does not suffice to explain users’ polarized news browsing behaviors. Rather, users’ explicit choices also contribute to the observed polarization. Our findings are thus overall in line with previous studies on the role of user choices (Bakshy, Messing, and Adamic 2015), while contributing results based on a novel dataset of unprecedented detail and from novel angles of analysis.

Limitations. Our work should be seen in the light of its limitations:

1. **Opt-in sample:** While we strongly believe the opt-in nature of our recruitment is necessary for an ethical research approach, it may well have provided a biased sample.
2. **Domain-level analysis:** We aggregated all pages in a domain when constructing the browsing, co-browsing, and hyperlink graphs, essentially treating all pages in a domain as interchangeable. This was done in order to avoid sparsity issues that would arise when aggregating at the URL level instead.
3. **Personalization:** Our analysis did not account for the role of personalization algorithms in browsing. While this is limiting to an extent, we point out that algorithmic personalization is still rare in online news (Thurman and Schiffrer 2012).
4. **Correlation vs. causation:** Although we showed that the polarization baked into the hyperlink structure of the Web alone does not suffice to explain users’ polarized news consumption patterns, we do not claim to have revealed

the causes of polarization in news consumption, which constitutes an important direction for future work.

Future work: the role of the center. Our analysis raises the profile of the role of the center as a clearly identifiable area in polarization studies, which proposes a potentially valuable role of centrist websites as a location for bipartisanship. These sites may themselves benefit from closer study and analysis, particularly when analyzed at the webpage or topic level. This becomes increasingly important when, as today, we see degrees of partisanship that are unprecedented and arguably unhealthy for society. Understanding where social worlds intersect provides opportunities for communication and potentially reconciliation within a polarized and separated populace. We should ask: Can centrist websites serve as a “demilitarized zone” of the Web, and can an analysis such as ours, but focused on centrist websites, provide implications for the improved design of virtual spaces?

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References

- Bakshy, E.; Messing, S.; and Adamic, L. A. 2015. Exposure to ideologically diverse news and opinion on Facebook. *Science* 348(6239): 1130–1132.
- Barberá, P.; Jost, J. T.; Nagler, J.; Tucker, J. A.; and Bonneau, R. 2015. Tweeting from left to right: Is online political communication more than an echo chamber? *Psychological Science* 26(10): 1531–1542.
- Barthel, M.; Shearer, E.; Gottfried, J.; and Mitchell, A. 2015. The evolving role of news on Twitter and Facebook. *Pew Research Center*.
- Batagelj, V.; and Zaversnik, M. 2003. An $O(m)$ algorithm for cores decomposition of networks. *arXiv preprint cs/0310049*.
- De Meo, P.; Ferrara, E.; Fiumara, G.; and Provetti, A. 2011. Generalized Louvain method for community detection in large networks. In *ICISDA*.
- Dhillon, I. S. 2001. Co-clustering documents and words using bipartite spectral graph partitioning. In *KDD*.
- Flaxman, S.; Goel, S.; and Rao, J. M. 2016. Filter bubbles, echo chambers, and online news consumption. *Public Opinion Quarterly* 80(S1): 298–320.
- Garimella, K.; De Francisci Morales, G.; Gionis, A.; and Mathioudakis, M. 2018. Political discourse on social media: Echo chambers, gatekeepers, and the price of bipartisanship. In *WWW*.
- Garrett, R. K. 2009. Echo chambers online? Politically motivated selective exposure among Internet news users. *Journal of Computer-Mediated Communication* 14(2): 265–285.
- Geiger, A. 2019. Key findings about the online news landscape in America. *Pew Research Center*.
- Gentzkow, M.; and Shapiro, J. M. 2011. Ideological segregation online and offline. *The Quarterly Journal of Economics* 126(4): 1799–1839.
- Guess, A. M. 2018. (Almost) Everything in Moderation: New Evidence on Americans’ Online Media Diets. *American Journal of Political Science*.
- Iyengar, S.; and Hahn, K. S. 2009. Red media, blue media: Evidence of ideological selectivity in media use. *Journal of Communication* 59(1): 19–39.
- Iyengar, S.; Lelkes, Y.; Levendusky, M.; Malhotra, N.; and Westwood, S. J. 2019. The origins and consequences of affective polarization in the United States. *Annual Review of Political Science* 22: 129–146.
- Klapper, J. T. 1960. *The effects of mass communication*. Free Press.
- Kulshrestha, J.; Eslami, M.; Messias, J.; Zafar, M. B.; Ghosh, S.; Gummadi, K. P.; and Karahalios, K. 2017. Quantifying search bias: Investigating sources of bias for political searches in social media. In *CSCW*.
- Lahoti, P.; Garimella, K.; and Gionis, A. 2018. Joint non-negative matrix factorization for learning ideological leaning on Twitter. In *WSDM*.
- Messing, S.; and Westwood, S. J. 2014. Selective exposure in the age of social media: Endorsements trump partisan source affiliation when selecting news online. *Communication Research* 41(8): 1042–1063.
- Meusel, R.; Vigna, S.; LehMBERG, O.; and Bizer, C. 2014. Graph structure in the Web—revisited: A trick of the heavy tail. In *WWW*.
- Narayanan, V.; Barash, V.; Kelly, J.; Kollanyi, B.; Neudert, L.-M.; and Howard, P. N. 2018. Polarization, partisanship and junk news consumption over social media in the US. *arXiv preprint arXiv:1803.01845*.
- Nelson, J. L.; and Webster, J. G. 2017. The myth of partisan selective exposure: A portrait of the online political news audience. *Social Media + Society* 3(3).
- Nikolov, D.; Lalmas, M.; Flammini, A.; and Menczer, F. 2019. Quantifying biases in online information exposure. *JAIST* 70(3): 218–229.
- Pariser, E. 2011. *The filter bubble: What the Internet is hiding from you*. Penguin UK.
- Parse.ly. 2016. Network Referrer Dashboard. Data retrieved from <https://www.parse.ly/resources/data-studies/referrer-dashboard>. Accessed: 2021-04-10.
- Peterson, E.; Goel, S.; and Iyengar, S. 2018. Echo chambers and partisan polarization: Evidence from the 2016 presidential campaign. *Political Science Research and Methods*.
- Raghavan, M.; Anderson, A.; and Kleinberg, J. 2018. Mapping the invocation structure of online political interaction. In *WWW*.

- Ribeiro, F. N.; Henrique, L.; Benevenuto, F.; Chakraborty, A.; Kulshrestha, J.; Babaei, M.; and Gummadi, K. P. 2018. Media bias monitor: Quantifying biases of social media news outlets at large-scale. In *ICWSM*.
- Robertson, R. E.; Jiang, S.; Joseph, K.; Friedland, L.; Lazer, D.; and Wilson, C. 2018. Auditing partisan audience bias within Google search. *Proceedings of the ACM on Human-Computer Interaction (CSCW) 2*: 1–22.
- Rogowski, J. C.; and Sutherland, J. L. 2016. How ideology fuels affective polarization. *Political Behavior 38*(2).
- Role, F.; Morbieu, S.; and Nadif, M. 2019. CoClust: A Python Package for Co-Clustering. *Journal of Statistical Software 88*(1): 1–29.
- Schmidt, A. L.; Zollo, F.; Del Vicario, M.; Bessi, A.; Scala, A.; Caldarelli, G.; Stanley, H. E.; and Quattrociocchi, W. 2017. Anatomy of news consumption on Facebook. *Proceedings of the National Academy of Sciences 114*(12): 3035–3039.
- Sears, D. O.; and Freedman, J. L. 1967. Selective exposure to information: A critical review. *Public Opinion Quarterly 31*(2): 194–213.
- Stroud, N. J. 2010. Polarization and partisan selective exposure. *Journal of Communication 60*(3): 556–576.
- Sunstein, C. R. 2009. *Republic.com 2.0*. Princeton University Press.
- Thurman, N.; and Schifferes, S. 2012. The future of personalization at news websites: Lessons from a longitudinal study. *Journalism Studies 13*(5-6): 775–790.
- Webster, S. W.; and Abramowitz, A. I. 2017. The ideological foundations of affective polarization in the US electorate. *American Politics Research 45*(4): 621–647.